The Econometrics of Matching

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Bringing Matching Theory to Data

Basic, one-dimensional theory yields stark predictions (assortative matching) that only approximately hold in the data In the data, almost every "matching cell" is non-empty. To reconcile theory and data :

- can add frictions (see Robin and Shimer)
- or assume a reduced-form relationship (the rank-order property of Fox)
- or model directly matching on both observables and unobservables (our main focus here.)

Main Themes

- More emphasis on TU than on NTU
- Focus on one-to-one rather than many-to-(one, many) (unlike Fox)
- These models are hard to identify, depending on what we observe/what we assume.

Matching assumed frictionless here : no search cost, perfect information among participants, transfers (if any) are costless.

Common Structures

- Populations : X et Y, measures F et G ("men" and "women")
- Matching : measure μ on $XY = (X \times Y) \cup (X \times 0) \cup (0 \times Y)$, given marginals F et G
- normalize utilities of singles to 0
- NTU : match (x, y) brings (algebraic, ordinal) gain U(x, y) to x and V(x, y) to y
- (perfectly) TU : it creates (algebraic) joint surplus s (x, y), endogenously shared between the partners.

Common equilibrium concept : stability

- no one is better off divorcing their current match and remaining single
- NTU : no (x, y) pair is better off divorcing their current matches and rematching together
- TU : no (*x*, *y*) pair can find a way of sharing *s*(*x*, *y*) that improves over both of their current outcomes.

Consequences for NTU : coupled equations

NTU : stable matchings solve

$$u(x) = \max_{z} \{ U(x,z) | V(x,z) \ge v(z) \}$$

and

$$v(y) = \max_{z} \{V(z, y) | U(z, y) \ge u(z)\}$$

for some pair of functions *u* and *v*.



TU : stable matchings solve

$$u(x) = \max_{z} \{s(x,z) - v(z)\}$$

and

$$v(y) = \max_{z} \{s(z, y) - u(z)\}$$

for some pair of functions *u* and *v*.

The Data

We observe a realized matching, and perhaps a noisy measure of outcomes (e.g. children, divorce, labor market dynamics...) We partition payoff-relevant types into observed (by the econometrician) and unobserved characteristics :

$$\mathbf{x} = (\mathbf{I}, \varepsilon), \ \mathbf{y} = (\mathbf{J}, \eta)$$

We only observe $\bar{\mu}(I, J) = E(\mu(x, y)|I, J)$ and perhaps (say in TU)

$$\bar{\mathbf{s}}(\mathbf{I},\mathbf{J})=E\left(\tilde{\mathbf{s}}(\mathbf{x},\mathbf{y})|\mathbf{I},\mathbf{J}\right)$$

with \tilde{s} a noisy measure of s (and $E \equiv E_{\mu}$.)

The Empirical Content of Matching Models

The theory implies the coupled equations, and only that : given *s*, *F*, *G*, and a solution (*u*, *v*), any μ

- whose support is contained in the arg max_z
- which integrates to the margins F and G
- is a stable matching.

Four approaches

- Choo-Siow : impose separability (unobserved characteristics do not interact)
- Chiappori-Oreffice-Quintana-Domeque : assume agreement on "attractiveness indices" (lower-dimensional matching)
- Fox : assume a rank-order property
- Hitsch-Hortacsu-Ariely : build on the coupled equations.

Conditional just-identification in TU models

Choo-Siow (2006, multinomial logit) + Galichon-Salanié (2011, general case) : even if we assume separability (S) : $s(x, y) = \bar{s}(I, J) + \varepsilon_I(J) + \eta_J(I)$, **and** we (assume that we) know the distributions \mathcal{P}_I of $\varepsilon_I(.)$ and Q_J of $\eta_J(.)$, the function \bar{s} is just-identified from data on the observed matching $\bar{\mu}$ on one cross-section.

=absence of complementarities on unobserved characteristics : if

$$\mathbf{x} = (\mathbf{I}, \varepsilon); \ \mathbf{x}' = (\mathbf{I}, \varepsilon')$$

and

$$\mathbf{y} = (\mathbf{J}, \eta); \ \mathbf{y}' = (\mathbf{J}, \eta')$$

then s(x, y) + s(x', y') = s(x, y') + s(x', y). How strong ? depends on data, and on question asked from it.

Separability buys us discrete choice

Then

$$u(x) = \max_{z} \left(s(x,z) - v(z) \right)$$

gives

$$u(x) = \max_{J} \left(\bar{s}(I, J) + \varepsilon_{I}(J) - \min_{\eta} \left(v(J, \eta) - \eta_{J}(I) \right) \right)$$

Denote V(I, J) the min_{η} and $U(I, J) = \bar{s}(I, J) - V(I, J)$;

$$u(x) = \max_{J} \left(U(I, J) + \varepsilon_{I}(J) \right).$$

Just-identification

See Salanié, tomorrow : in a large market, if we "know" the \mathcal{P}_l and Q_J , we can compute easily

$$\bar{s}(I, J) = ...; u, v, U, V = ...$$
 etc.

Large market : large numbers of individuals for any observed type

Ensures proportions \equiv probabilities, and uniqueness of market equilibrium payoffs *u* and *v*.

Beyond just-identification : testing

Just identification rules out testing, unless

$$\bar{s}(I,J) = \sum_{k=1}^{K} \lambda^k \bar{s}_k(I,J)$$

for some unknown λ 's and known basis functions \bar{s}_k .

we pool data on several "markets" and restrict s

 i, *J*) and/or the variation of the P₁ and Q_J across markets. E.g., Chiappori-Salanié-Weiss : stability of complementarities and heteroskedasticity across cohorts.

Beyond just-identification : the distribution of errors

The \mathcal{P}_{l} and Q_{J} have many degrees of freedom; how can we identify some of them/test the model? again : restrict the specification of the mean surplus, and/or pool data across markets.

Estimation in the separable framework

First estimate the margins on observed characteristics : $\hat{F}(I), \hat{G}(J)$.

If the \mathcal{P}_I and Q_J are fully specified and the $\bar{s}(I, J)$ is

unrestricted : apply the closed-form formula

If they depend on a parameter vector λ : can use the formula as the basis of a minimum-distance estimator,

or do constrained maximum-likelihood estimation.

Constrained MLE

$$\max_{\lambda} \left(2\sum_{IJ} \hat{\mu}_{IJ} \ln \mu_{IJ}^{\lambda} + \sum_{I} \hat{\mu}_{I0} \ln \mu_{I0}^{\lambda} + \sum_{J} \hat{\mu}_{0J} \ln \mu_{0J}^{\lambda} \right)$$

where μ^{λ} is the optimal matching for parameter vector λ . *Problem :* how do we compute μ^{λ} ?

Iterated Projection Fitting Procedure

Galichon-Salanié : the closed form formula links $\bar{s}^{\lambda}(I, J)$, μ_{IJ}^{λ} , μ_{I0}^{λ} , μ_{0J}^{λ} and λ .

The difficulty : make sure that μ^{λ} fits the margins F(I), G(J). The solution : start from a well-chosen $\mu^{\lambda,0}$, project on the margins for men (*F*), then on the margins for women (*G*), and iterate

using the proper quasi-distance (a Bregman divergence) dictated by the specification of the errors for λ .

A Two-index Model of Surplus

Chiappori-Oreffice-Quintana-Domeque assume that there exists two *scalar* functions m and f such that

$$s(I, J, \varepsilon, \eta) = S(m(I), \varepsilon, f(J), \eta)$$

and $\varepsilon \perp J | f(J), \eta \perp I | m(I)$.

e.g. m(I) is the socially agreed (by women) attractiveness index for men.

Then $\mu(I, J) \equiv \nu(m(I), f(J));$

and on realized matches : $J \perp I \mid m(I)$ and $I \perp J \mid f(J)$.

Estimation and Testing

Regressing observed characteristics of wife J on observed characteristics of husband I identifies the indifference curves of m(I);

cross-equation restrictions : same index *m* for *all* components of $J \rightarrow$ testable.

Generalizes to multi-index models.

Comparing Similar Markets

Suppose we observe many markets that

- are separated (no matching across markets)
- have the same surplus function s(x, y)
- but differ in their sample margins.

Each such market *n* has a list of observed characteristics C_n (number of men of type I = 1, of women of type J = 3...) and an associated matching μ_n (how many such men marry such women...)

and we can estimate $Pr(\mu_n = \mu | C_n = C.)$

Fox's Rank-order Property

States that for every *C* and all μ^1 , μ^2 feasible for *C*

$$\Pr(\mu_n = \mu^1 | C_n = C) \ge \Pr(\mu_n = \mu^2 | C_n = C) \text{ iff } \sum_{\mu^1} \bar{s}(I, J) \ge \sum_{\mu^2} \bar{s}(I, J).$$

Intuition : in single-agent choice problems, under weak conditions $Pr(d = d_k | U_1, ..., U_K)$ is an increasing function of mean utilities U_k .

Pluses and Minuses

Maximum score estimation can rely on selected stability conditions

is easy to reformulate beyond one-to-one matching (Fox Bajari 2009, Fox 2011)

But not founded on a microeconomic model of matching; e.g. from Galichon-Salanié, with separability μ_n maximizes

$$\sum_{\mu} \bar{\mathbf{s}}(\mathbf{I}, \mathbf{J}) + \mathcal{E}(\mu; \mathbf{C}_n)$$

with \mathcal{E} a generalized entropy term that comes from non-random matching over unobservables.

NTU version

Hitsch-Hortacscu-Ariely (AER 2010) : man x is matched with woman y iff

$$U(x,y) \ge U(x,z) \quad \forall z \text{ s.t. } V(x,z) \ge v(y).$$

Specify $U(x, y) = \overline{U}(I, J; \lambda) + \xi_{xy}$, and v(y) as a fixed effect; if ξ_{xy} is iid type I EV, gives a conditional logit model to recover λ . Very restrictive assumption (more than separability; no (x, J)interaction for instance.)

Solving the System

Brute-force approach : specify $U(x, y; \lambda)$ and $V(x, y; \lambda)$, solve for $u(x; \lambda)$ and $v(y; \lambda)$, get $\mu(x, y; \lambda)$ and take it to the data. (TU version : only specify $s(x, y; \lambda)$.) In general, using the coupled equations is computationally intensive And identification is mostly unexplored.

Concluding remarks

A lot remains to be done :

applications

testing of identifying assumptions in the various approaches bridging theory and data for "matching in contracts" frictions and dynamics (as in the job search literature.)